

PhD Defense
On the tradeoffs of statistical learning with privacy

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Jury Members

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4 Context

Increasing data usage :

- ▶ Natural Language Processing (ChatGPT, ...)
- ▶ Medical applications
- ▶ ...

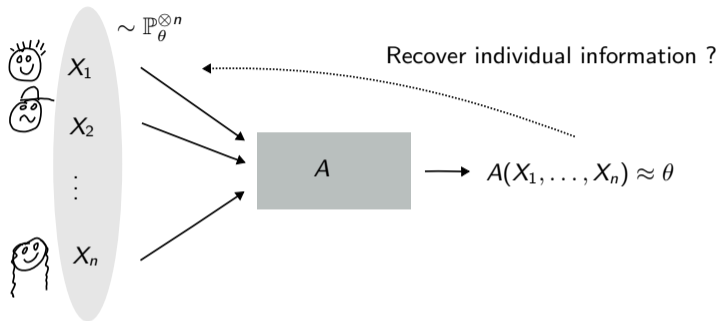
Observation :

Most applications are not interested in the dataset itself, but rather in some quantities defined at the scale of the population.

Threats to privacy :

Incautious use of data can leak personal information.

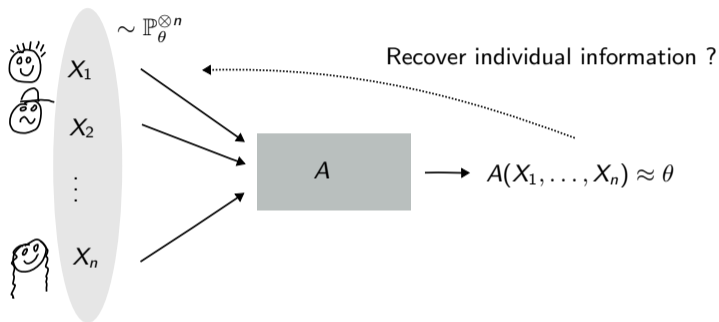
5 Context



Examples :

- ▶ Bernoulli estimation (proportion)
- ▶ Complex distributions
- ▶ ...

6 Context



Question :

*Is it possible to defend against **ANY** attack ?*

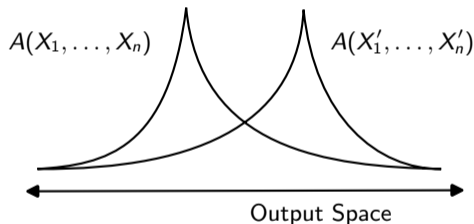
²Gonon et al., *Sparsity in neural networks can improve their privacy.*

7 Differential Privacy

Neighboring relation : $\mathbf{X} \sim \mathbf{X}'$ iff \mathbf{X} can be obtained from \mathbf{X}' by changing the data of one individual.

Definition (informal) : A is $\epsilon > 0$ -DP if $A(\mathbf{X})$ is ϵ -close to $A(\mathbf{X}')$ for any $\mathbf{X} \sim \mathbf{X}'$.³

Impact of ϵ : The smaller ϵ , the more privacy.



Question :

How does privacy affect classical statistical estimation results ?

³Cynthia Dwork et al. "Calibrating Noise to Sensitivity in Private Data Analysis". In: 2006.

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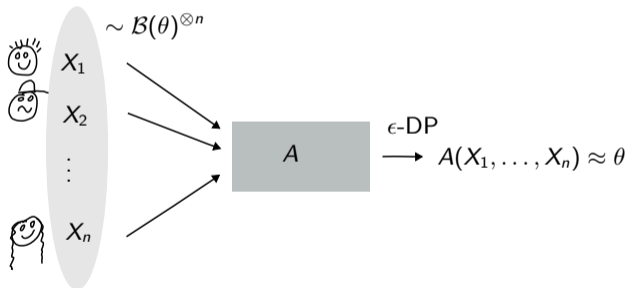
Quantiles Estimation

Conclusion

9 Bernoulli estimation setup

Setup : $\theta \in [0, 1]$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathcal{B}(\theta)^{\otimes n}$.

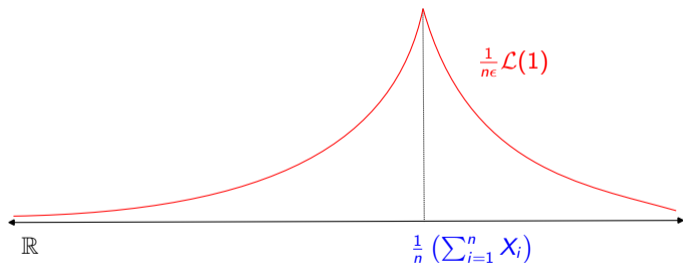
Measure of performance : $\mathbb{E}_{A, \mathbf{X}} ((A(\mathbf{X}) - \theta))^2$.



10 Private estimator

Laplace mechanism :

$$A(\mathbf{X}) = \frac{1}{n} \left(\sum_{i=1}^n X_i \right) + \frac{1}{n\epsilon} \mathcal{L}(1)$$



11 Error decomposition

Laplace mechanism :

$$A(\mathbf{X}) = \frac{1}{n} \left(\sum_{i=1}^n X_i \right) + \frac{1}{n\epsilon} \mathcal{L}(1)$$

is ϵ -DP.

Error :

$$\mathbb{E}_{\mathbb{P}_A, \mathcal{B}(\theta)^{\otimes n}} \left((A(\mathbf{X}) - \theta)^2 \right) \leq \frac{1/4}{n} + \frac{2}{n^2\epsilon^2}$$

Two regimes :

- ▶ **Low privacy regime :** $\epsilon = \Omega(1/\sqrt{n})$, no significant effect on estimation.
- ▶ **High privacy regime :** $\epsilon \ll 1/\sqrt{n}$, the precision can be arbitrarily degraded.

Question :

Is it possible to do better ?

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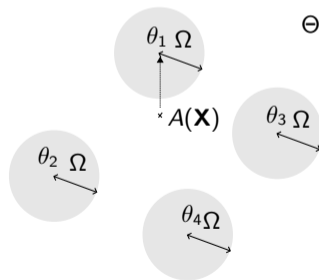
Conclusion

13 Minimax risk and reduction to hypothesis testing

Setup : $\theta \in \Theta$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\theta^{\otimes n}$.

Minimax risk :⁴⁵

$$\begin{aligned} \mathfrak{M}_n &:= \inf_A \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{X} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \inf_A \sup_{i=1, \dots, N} \mathbb{E}_{A, \mathbf{X} \sim \theta_i} (\text{Error}(A(\mathbf{X}), \theta_i)) \\ &\geq \Phi(\Omega) \sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \end{aligned}$$



⁴Alexandre B. Tsybakov. *Introduction to Nonparametric Estimation*. 2009.

⁵ $\text{Error}(\cdot, \cdot) = \Phi(d(\cdot, \cdot))$

14 Without taking privacy into account

Setup : $\theta \in \Theta$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\theta^{\otimes n}$.

Minimax risk :

$$\begin{aligned} \mathfrak{M}_n &:= \inf_A \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{X} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \inf_A \sup_{i=1, \dots, N} \mathbb{E}_{A, \mathbf{X} \sim \theta_i} (\text{Error}(A(\mathbf{X}), \theta_i)) \\ &\geq \phi(\Omega) \sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \end{aligned}$$

Le Cam's lemma :⁶ $\sup_{i=1,2} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \geq \frac{1}{2} (1 - \text{TV}(\mathbb{P}_{\theta_1}^{\otimes n}, \mathbb{P}_{\theta_2}^{\otimes n}))$.

Fano's lemma :⁷ For any dominating measure \mathbb{Q} ,

$$\sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \geq 1 - \frac{1 + \frac{1}{N} \sum_{i=1}^N \text{KL}(\mathbb{P}_{\theta_i}^{\otimes n} \parallel \mathbb{Q})}{\ln(N)}.$$

⁶ $\text{TV}(\mathbb{P}_1, \mathbb{P}_2) := \sup_S \mathbb{P}_1(S) - \mathbb{P}_2(S)$

⁷ $\text{KL}(\mathbb{P}_1 \parallel \mathbb{P}_2) := \int \ln \left(\frac{d\mathbb{P}_1}{d\mathbb{P}_2} \right) d\mathbb{P}_1$

15 Without taking privacy into account

Setup : $\theta \in \Theta$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\theta^{\otimes n}$.

Minimax risk :

$$\begin{aligned} \mathfrak{M}_n &:= \inf_A \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{X} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \inf_A \sup_{i=1, \dots, N} \mathbb{E}_{A, \mathbf{X} \sim \theta_i} (\text{Error}(A(\mathbf{X}), \theta_i)) \\ &\geq \phi(\Omega) \sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \end{aligned}$$

Question :

Is it possible to obtain similar lower-bounds that take privacy into consideration ?

16 Reduction to a transport problem

Setup : $\theta \in \Theta$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\theta^{\otimes n}$.

Minimax risk :

$$\begin{aligned} \mathfrak{M}_n &:= \inf_A \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{X} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \inf_A \sup_{i=1, \dots, N} \mathbb{E}_{A, \mathbf{X} \sim \theta_i} (\text{Error}(A(\mathbf{X}), \theta_i)) \\ &\geq \Phi(\Omega) \sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} \left(\hat{i}(A(\mathbf{X})) \neq i \right) \end{aligned}$$

Reduction to a transport problem :⁸

$$\sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{X} \sim \theta_i} \left(\hat{i}(A(\mathbf{X})) \neq i \right) \geq \boxed{\sup_{\mathbb{Q} \in \Pi(\mathbb{P}_1^{\otimes n}, \dots, \mathbb{P}_N^{\otimes n})} \int s(\mathbf{X}_1, \dots, \mathbf{X}_N) d\mathbb{Q}(\mathbf{X}_1, \dots, \mathbf{X}_N)},$$

where s is a *similarity function* satisfying, for any $\mathbf{X}_1, \dots, \mathbf{X}_N$,

$$\boxed{\frac{1}{N} \sum_{i=1}^N \mathbb{P}_A \left(\hat{i}(A(\mathbf{X}_i)) \neq i \right) \geq s(\mathbf{X}_1, \dots, \mathbf{X}_N)}.$$

⁸Clément Lalanne, Aurélien Garivier, and Rémi Gribonval. "On the Statistical Complexity of Estimation and Testing under Privacy Constraints". In: (2023).

17 Building similarity functions

$$\frac{1}{N} \sum_{i=1}^N \mathbb{P}_A \left(\hat{i}(A(\mathbf{X}_i)) \neq i \right) \geq s(\mathbf{X}_1, \dots, \mathbf{X}_N).$$

Definition : A is ϵ -DP if $\mathbf{X} \sim \mathbf{X}' \implies \mathbb{P}(A(\mathbf{X}) \in S) \leq e^\epsilon \times \mathbb{P}(A(\mathbf{X}') \in S)$.⁹

Two marginals : $\frac{1}{2} \sum_{i=1}^2 \mathbb{P}_A \left(\hat{i}(A(\mathbf{X}_i)) \neq i \right) \geq \frac{1}{2} e^{-\epsilon d_{\text{ham}}(\mathbf{X}_1, \mathbf{X}_2)}$ ¹⁰

Many marginals : $\frac{1}{N} \sum_{i=1}^N \mathbb{P}_A \left(\hat{i}(A(\mathbf{X}_i)) \neq i \right) \geq 1 - \frac{1 + \frac{\epsilon}{N^2} \sum_{i=1}^N \sum_{j=1}^N d_{\text{ham}}(\mathbf{X}_i, \mathbf{X}_j)}{\ln(N)}$

⁹Dwork et al., "Calibrating Noise to Sensitivity in Private Data Analysis".

¹⁰ $d_{\text{ham}}((X_1, \dots, X_n), (X'_1, \dots, X'_n)) := \sum_{i=1}^n \mathbb{1}_{X_i \neq X'_i}$

18 Back to the transport problem

$$\sup_{\mathbb{Q} \in \Pi(\mathbb{P}_1^{\otimes n}, \dots, \mathbb{P}_N^{\otimes n})} \int s(\mathbf{X}_1, \dots, \mathbf{X}_N) d\mathbb{Q}(\mathbf{X}_1, \dots, \mathbf{X}_N),$$

where s is **non-increasing** in $d_{\text{ham}}(\mathbf{X}_i, \mathbf{X}_j)$ for any i, j .

Question :

How to construct a coupling that makes those quantities big ?

19 A good enough coupling

$$\sup_{\mathbb{Q} \in \Pi(\mathbb{P}_1^{\otimes n}, \dots, \mathbb{P}_N^{\otimes n})} \int s(\mathbf{X}_1, \dots, \mathbf{X}_N) d\mathbb{Q}(\mathbf{X}_1, \dots, \mathbf{X}_N),$$

where s is **non-increasing** in $d_{\text{ham}}(\mathbf{X}_i, \mathbf{X}_j)$ for any i, j .

Near optimal coupling for equalities : There exists $(X_i)_{i=1, \dots, N}$ of distribution χ , a coupling between $(\mathbb{P}_i)_{i=1, \dots, N}$ such that¹¹

$$\forall i, j, \quad \text{TV}(\mathbb{P}_i, \mathbb{P}_j) \leq \mathbb{P}(X_i \neq X_j) \leq \frac{2\text{TV}(\mathbb{P}_i, \mathbb{P}_j)}{1 + \text{TV}(\mathbb{P}_i, \mathbb{P}_j)}.$$

Final coupling : $\mathbb{Q}^* = \chi^{\otimes n}$

¹¹Omer Angel and Yinon Spinka. *Pairwise optimal coupling of multiple random variables*. 2021.

20 Final results

Setup : $\theta \in \Theta$, $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\theta^{\otimes n}$.

Minimax risk :

$$\begin{aligned} \mathfrak{M}_n &:= \inf_A \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{x} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \inf_A \sup_{i=1, \dots, N} \mathbb{E}_{A, \mathbf{x} \sim \theta_i} (\text{Error}(A(\mathbf{X}), \theta_i)) \\ &\geq \phi(\Omega) \sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{x} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \end{aligned}$$

Private Le Cam's lemma :¹²

$$\sup_{i=1,2} \mathbb{P}_{A, \mathbf{x} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \geq \frac{1}{2} (1 - (1 - e^{-\epsilon}) \text{TV}(\mathbb{P}_{\theta_1}, \mathbb{P}_{\theta_2}))^n .$$

Private Fano's lemma :¹³ For any dominating measure \mathbb{Q} ,

$$\sup_{i=1, \dots, N} \mathbb{P}_{A, \mathbf{x} \sim \theta_i} (\hat{i}(A(\mathbf{X})) \neq i) \geq 1 - \frac{1 + \frac{n\epsilon}{N^2} \sum_{i,j=1}^N \frac{2\text{TV}(\mathbb{P}_{\theta_i}, \mathbb{P}_{\theta_j})}{1 + \text{TV}(\mathbb{P}_{\theta_i}, \mathbb{P}_{\theta_j})}}{\ln(N)} .$$

¹² $\text{TV}(\mathbb{P}_1, \mathbb{P}_2) := \sup_S \mathbb{P}_1(S) - \mathbb{P}_2(S)$

¹³ $\text{KL}(\mathbb{P}_1 \parallel \mathbb{P}_2) := \int \ln \left(\frac{d\mathbb{P}_1}{d\mathbb{P}_2} \right) d\mathbb{P}_1$

21 Illustration

$$\inf_{A} \sup_{\theta \in \Theta} \mathbb{E}_{A, \mathbf{X} \sim \theta} (\text{Error}(A(\mathbf{X}), \theta)) \geq \phi(\Omega) \left(1 - \frac{1 + \frac{n\epsilon}{N^2} \sum_{i,j=1}^N \frac{2\text{TV}(\mathbb{P}_{\theta_i}, \mathbb{P}_{\theta_j})}{1 + \text{TV}(\mathbb{P}_{\theta}, \mathbb{P}_{\theta_j})}}{\ln(N)} \right)$$

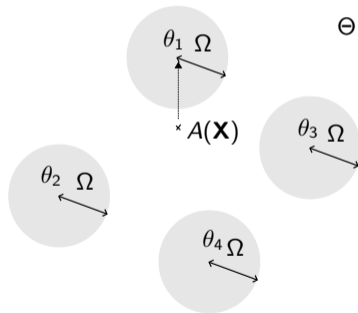


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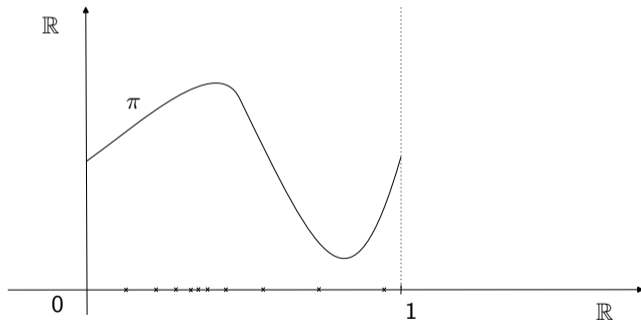
Quantiles Estimation

Conclusion

23 Density Estimation Problem

Setup : $\mathbf{X} = (X_1, \dots, X_n) \sim \mathbb{P}_\pi^{\otimes n}$, π a **density of probability** with respect to Lebesgue's measure on $[0, 1]$.

Measure of performance : $\mathbb{E} (\|A(\mathbf{X}) - \pi\|_{L_2}^2)$.



24 L_2 approximations and projection estimators

Reference Fourier basis :

$$\begin{aligned}\phi_1(x) &= 1 \\ \phi_{2k}(x) &= \sqrt{2} \sin(2\pi kx) \quad k \geq 1 \\ \phi_{2k+1}(x) &= \sqrt{2} \cos(2\pi kx) \quad k \geq 1.\end{aligned}$$

L_2 approximation :

$$\sum_{i=1}^N \theta_i \phi_i \xrightarrow[N \rightarrow +\infty]{L^2} \pi \quad \text{where} \quad \theta_i := \int_{[0,1]} \pi \phi_i.$$

Projection estimator :¹⁴

$$\hat{\pi}^{\text{proj}}(\mathbf{X}) = \sum_{i=1}^N \hat{\theta}_i \phi_i \quad \text{where} \quad \hat{\theta}_i := \frac{1}{n} \sum_{j=1}^n \phi_i(X_j).$$

Question : *How do we add privacy ?*

¹⁴Tsybakov, *Introduction to Nonparametric Estimation*.

25 L_2 approximations and *private* projection estimators

Private projection estimator :¹⁵

$$\hat{\pi}^{\text{proj}}(\mathbf{X}) = \sum_{i=1}^N \left(\hat{\theta}_i + C_{\epsilon, N} \mathcal{L}(1) \right) \phi_i \quad \text{where} \quad \hat{\theta}_i := \frac{1}{n} \sum_{j=1}^n \phi_i(X_j) .$$

$\hat{\pi}^{\text{proj}}$ is ϵ -DP.

Question :

What is the utility (error) of this estimator ?

¹⁵Larry A. Wasserman and Shuheng Zhou. "A Statistical Framework for Differential Privacy". In: (2010).

26 Sobolev spaces and approximation speed

$$\hat{\pi}^{\text{proj}}(\mathbf{X}) = \sum_{i=1}^N \left(\hat{\theta}_i + C_{\epsilon, N} \mathcal{L}(1) \right) \phi_i \quad \text{where} \quad \hat{\theta}_i := \frac{1}{n} \sum_{j=1}^n \phi_i(X_j).$$

Sobolev spaces :

$$\Theta_{L, \beta}^{\text{PSob}} := \left\{ \pi \mid \int_{[0,1]} \left(\pi^{(\beta)} \right)^2 \leq L^2 \text{ and a few other minor hypotheses} \right\}. \quad (1)$$

Approximation speed :¹⁶ When $\pi \in \Theta_{L, \beta}^{\text{PSob}}$,

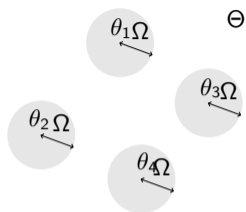
$$\mathbb{E} \left(\|\hat{\pi}^{\text{proj}}(\mathbf{X}) - \pi\|_{L_2}^2 \right) \leq C_{L, \beta} \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\epsilon)^{-\frac{2\beta}{\beta+3/2}} \right\}$$

Question :

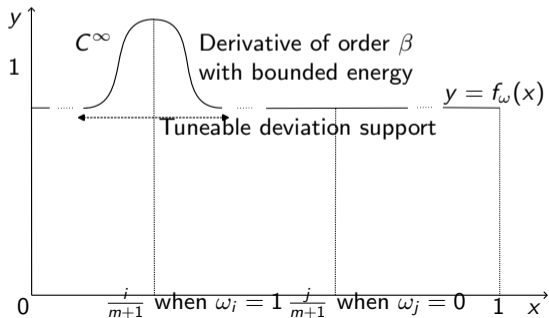
What about lower-bounds ?

¹⁶Clément Lalanne, Aurélien Garivier, and Rémi Gribonval. "About the Cost of Central Privacy in Density Estimation". In: (2023).

27 Lower-bounds



A packing of densities $f_{\omega_1}, f_{\omega_2}, \dots$ where for any $\omega \in \{0, 1\}^m$,



28 Lower-bounds

Lower-bound against ϵ -DP estimators :¹⁷

$$\inf_A \sup_{\pi} \mathbb{E} \left(\|A(\mathbf{X}) - \pi\|_{L_2}^2 \right) \geq C_{L,\beta} \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\epsilon)^{-\frac{2\beta}{\beta+1}} \right\}$$

Best known upper-bound :

$$\mathbb{E} \left(\|\hat{\pi}^{\text{proj}}(\mathbf{X}) - \pi\|_{L_2}^2 \right) \leq C_{L,\beta} \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\epsilon)^{-\frac{2\beta}{\beta+3/2}} \right\}$$

Question :

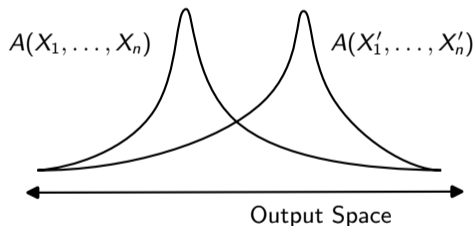
Is it possible to bridge the gap ?

¹⁷Lalanne, Garivier, and Gribonval, "About the Cost of Central Privacy in Density Estimation".

29 Concentrated Differential Privacy

Definition :¹⁸¹⁹ A is ρ -zCDP if $\mathbf{X} \sim \mathbf{X}' \implies \forall \alpha > 0, D_\alpha (A(\mathbf{X}) \| A(\mathbf{X}')) \leq \alpha \rho$, where

$$D_\alpha (\mathbb{P} \| \mathbb{Q}) := \frac{1}{\alpha - 1} \ln \int \left(\frac{d\mathbb{P}}{d\mathbb{Q}} \right)^{\alpha - 1} d\mathbb{Q}.$$



¹⁸Cynthia Dwork and Guy N Rothblum. "Concentrated differential privacy". In: (2016).

¹⁹Mark Bun and Thomas Steinke. "Concentrated Differential Privacy: Simplifications, Extensions, and Lower Bounds". In: 2016.

30 Lower-bounds and upper-bounds with CDP

Private projection estimator :

$$\hat{\pi}^{\text{proj}}(\mathbf{X}) = \sum_{i=1}^N \left(\hat{\theta}_i + C_{\rho, N} \mathcal{N}(0, 1) \right) \phi_i \quad \text{where} \quad \hat{\theta}_i := \frac{1}{n} \sum_{j=1}^n \phi_i(X_j) .$$

$\hat{\pi}^{\text{proj}}$ is ρ -zCDP.

Resulting upper-bound :

$$\mathbb{E} \left(\|\hat{\pi}^{\text{proj}}(\mathbf{X}) - \pi\|_{L_2}^2 \right) \leq C_{L, \beta} \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\sqrt{\rho})^{-\frac{2\beta}{\beta+1}} \right\}$$

Lower-bound against ρ -zCDP estimators :

$$\inf_A \sup_{\pi} \mathbb{E} \left(\|A(\mathbf{X}) - \pi\|_{L_2}^2 \right) \geq C_{L, \beta} \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\sqrt{\rho})^{-\frac{2\beta}{\beta+1}} \right\}$$

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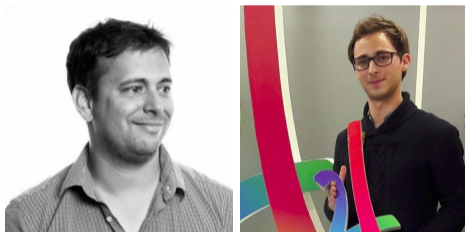
Density Estimation

Quantiles Estimation

Conclusion

32 Joint work with

Nicolas Grislain and Clément Gastaud from Sarus Technologies²⁰

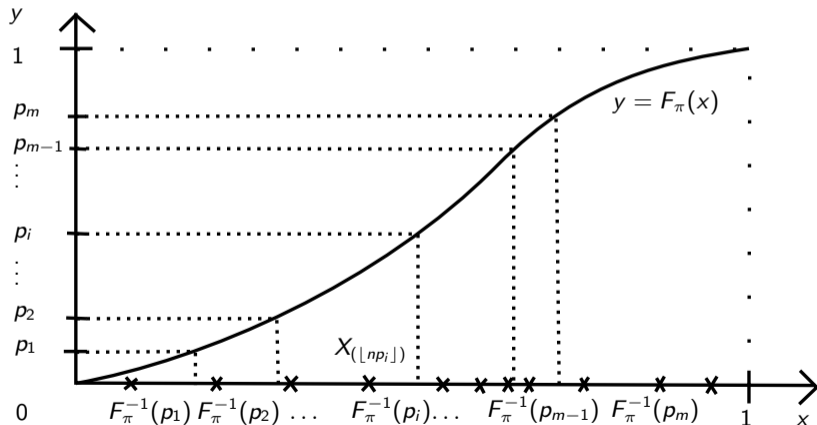


²⁰Clément Lalanne et al. "Private quantiles estimation in the presence of atoms". In: (2023). ISSN: 2049-8772.

33 Quantiles Estimation Problem

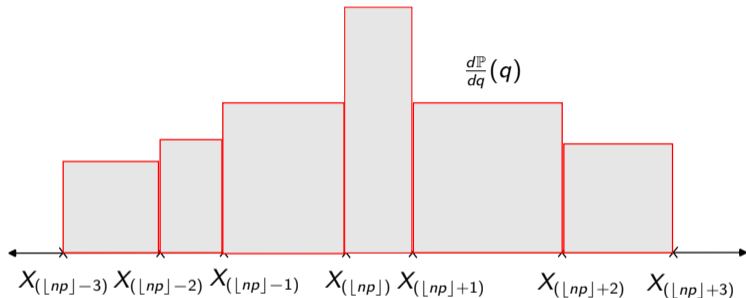
Inputs : samples $\mathbf{X} = (X_1, \dots, X_n) \stackrel{\text{i.i.d.}}{\sim} \mathbb{P}_\pi$ $\mathbf{p} = (p_1, \dots, p_m) \in (0, 1)^m$ sorted.

Desired output : Quantile estimator $\mathbf{q} \in [0, 1]^m$ of $(F_\pi^{-1}(p_1), \dots, F_\pi^{-1}(p_m))$.



34 Private Exponential Quantiles

Mechanism :²¹ For a single quantile q (associated with p),



Concentration result :²³ When π is away from 0 on a neighborhood of $F_\pi^{-1}(p)$,

$$\mathbb{P}(|q - F_\pi^{-1}(p)| > \gamma) \leq P(n) \max \left(e^{-C_1 \epsilon n \gamma}, e^{-C_2 \gamma^2 n} \right).$$

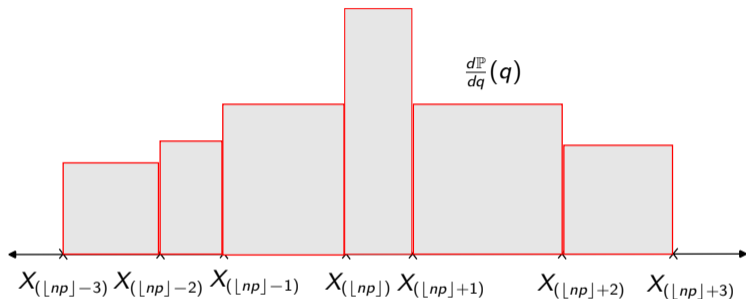
²¹Adam D. Smith. "Privacy-preserving statistical estimation with optimal convergence rates". In: 2011.

²² $d\mathbb{P}(q) \propto e^{-\frac{\epsilon}{2} |\{i | X_i < q\} - \lfloor np \rfloor|} dq$

²³Clément Lalanne, Aurélien Garivier, and Rémi Gribonval. "Private Statistical Estimation of Many Quantiles". In: 2023.

Independent Private Quantiles

Idea : Use QExp independently on \mathbf{p} with simple composition.



Concentration result :²⁵ When π is away from 0 on a neighborhood of $F_{\pi}^{-1}(\mathbf{p})$,

$$\mathbb{P}\left(\|\mathbf{q} - F_{\pi}^{-1}(\mathbf{p})\|_{\infty} > \gamma\right) \leq P(n, m) \max\left(e^{-C_1 \frac{\epsilon n \gamma}{m}}, e^{-C_2 \gamma^2 n}\right).$$

²⁴ $d\mathbb{P}(q_i) \propto e^{-\frac{\epsilon}{2m} |\{i | X_i < q_i\} - \lfloor np_i \rfloor|} dq_i$

²⁵ Lalanne, Garivier, and Gribonval, "Private Statistical Estimation of Many Quantiles".

36 Joint Exponential Private Quantiles

Idea :²⁶ Leverage structural dependencies. Quantiles are non-decreasing, between q_i and q_j should fall approximately $n|p_i - p_j|$ points. ²⁷

"Fun" discovery :²⁸ JointExp \approx Inverse Sensitivity Mechanism²⁹.

Consistency result : When π is away from 0 on a neighborhood of $F_\pi^{-1}(\mathbf{p})$, JointExp is consistent.

²⁶Jennifer Gillenwater, Matthew Joseph, and Alex Kulesza. "Differentially Private Quantiles". In: 2021.

²⁷ $d\mathbb{P}(\mathbf{q}) \propto e^{-\frac{\epsilon}{2} \sum_{i=1}^{m+1} |\delta^{\text{JE}}(i, \mathbf{X}, \mathbf{q})|} d\mathbf{q}$ where $\delta^{\text{JE}}(i, \mathbf{X}, \mathbf{q}) := n(p_i - p_{i-1}) - \#(\mathbf{X} \cap (q_{i-1}, q_i])$

²⁸Lalanne et al., "Private quantiles estimation in the presence of atoms".

²⁹Hilal Asi and John C. Duchi. "Near Instance-Optimality in Differential Privacy". In: CoRR (2020). arXiv: 2005.10630.

37 Recursive Private Quantiles

Idea :³⁰ Use QExp recursively with a dichotomy on \mathbf{p} .

Concentration result :³¹ When π is away from 0 on a neighborhood of $F_{\pi}^{-1}(\mathbf{p})$,

$$\mathbb{P}\left(\|\mathbf{q} - F_{\pi}^{-1}(\mathbf{p})\|_{\infty} > \gamma\right) \leq P(n, m) \max\left(e^{-C_1 \frac{\epsilon n \gamma}{(\log_2(m))^2}}, e^{-C_2 \gamma^2 n}\right).$$

Remark : Almost polylogarithmic degradation in m !

³⁰Haim Kaplan, Shachar Schnapp, and Uri Stemmer. "Differentially Private Approximate Quantiles". In: ed. by Kamalika Chaudhuri et al. PMLR, 2022.

³¹Lalanne, Garivier, and Gribonval, "Private Statistical Estimation of Many Quantiles".

38 Histograms

Idea :³²

$$\hat{\pi}^{\text{hist}}(t) := \sum_{b \in \text{bins}} \mathbb{1}_b(t) \frac{1}{nh} \left(\sum_{i=1}^n \mathbb{1}_b(X_i) + \frac{2}{\epsilon} \mathcal{L}_b \right).$$

Concentration result :³³ Bins of size h , $\gamma > C_4 h$, π is L -Lipschitz, I is a strict sub-interval

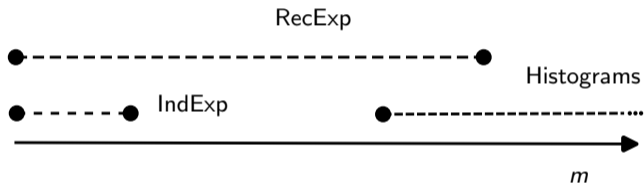
$$\begin{aligned} & \mathbb{P} \left(\|F_{\hat{\pi}^{\text{hist}}}^{-1} - F_{\pi}^{-1}\|_{\infty, I} > \gamma \right) \\ & \leq \frac{1}{h} e^{-C_1 \gamma h n \epsilon} + \frac{2}{h} e^{-C_2 h^2 (C_3 \gamma - L h)^2 n}. \end{aligned}$$

Remark : No degradation in m , but high entry cost.

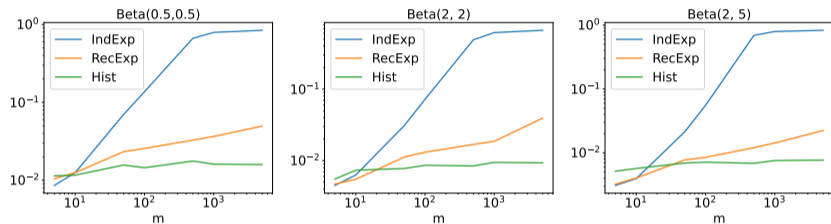
³²Wasserman and Zhou, "A Statistical Framework for Differential Privacy".

³³Lalanne, Garivier, and Gribonval, "Private Statistical Estimation of Many Quantiles".

39 Theoretical choice of algorithm



40 Numerical validation

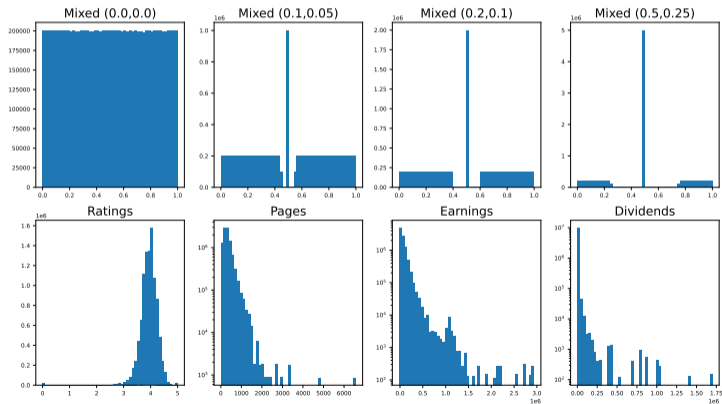


The vertical axis reads the error $\mathbb{E} (\|\mathbf{q} - F^{-1}(\mathbf{p})\|_{\infty})$ where $\mathbf{p} = \left(\frac{1}{4} + \frac{1}{2(m+1)}, \dots, \frac{1}{4} + \frac{m}{2(m+1)} \right)$ for different values of m , $n = 10000$, $\epsilon = 0.1$, and \mathbb{E} is estimated by Monte-Carlo averaging over 50 runs. The histogram is computed on 200 bins.

41 Dealing with atomic distributions

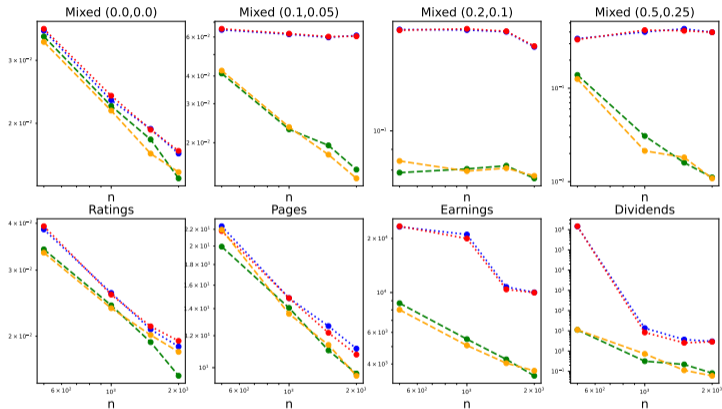
Inconsistency result :³⁴ When dealing with atomic distributions, all the *Exp mechanisms are inconsistent or have poor performances.

Proposed solution : Smoothing the distribution with noise addition can make those mechanisms consistent and helps the performances.



³⁴Lalanne et al., "Private quantiles estimation in the presence of atoms".

42 Dealing with atomic distributions



The vertical axis reads the error $\mathbb{E}(\|\hat{\mathbf{q}} - F^{-1}(\mathbf{p})\|_\infty)$ where $\mathbf{p} = \left(\frac{1}{m+1}, \dots, \frac{m}{m+1}\right)$ for $m = 8$, $\epsilon = 1$, $\hat{\mathbf{q}}$ is the private estimator, and \mathbb{E} is estimated by Monte-Carlo averaging over 50 runs.

Table of Contents

Introduction

An example : Bernoulli estimation

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44 Main Contributions of the Thesis and Take Home Messages

Depending on the level of privacy, the effects on the estimation can be either negligible, or they can be made arbitrarily bad.

On a positive note, there exist regimes of increasing privacy at a negligible cost.

The complexity of the statistical problem greatly affects those different regimes.

Main contributions :

- ▶ Propose a framework for lower-bounds that builds on recent coupling constructions.
- ▶ Give a unified view on the cost of privacy for multiple definitions of privacy.
- ▶ Advances on the private density estimation problem with matching or near-matching lower and upper-bounds.
- ▶ An in depth analysis of some of the statistical properties of existing mechanisms for the pointwise estimation of the quantile function.

Better similarity functions for different regimes :

- ▶ What we used : $\text{KL} (A(\mathbf{X}) \| A(\mathbf{Y})) \leq \epsilon d_{\text{ham}} (\mathbf{X}, \mathbf{Y})$
- ▶ State of the art³⁵ : $\text{KL} (A(\mathbf{X}) \| A(\mathbf{Y})) \leq \epsilon d_{\text{ham}} (\mathbf{X}, \mathbf{Y}) \frac{e^{\epsilon d_{\text{ham}}(\mathbf{X}, \mathbf{Y})} - 1}{e^{\epsilon d_{\text{ham}}(\mathbf{X}, \mathbf{Y})} + 1}$

Other constraints : Can it be applied to other forms of constraints (quantized, ...) ?

Numerical optimization :

$$\sup_{\mathbb{Q} \in \Pi(\mathbb{P}_1^{\otimes n}, \dots, \mathbb{P}_N^{\otimes n})} \int s(\mathbf{X}_1, \dots, \mathbf{X}_N) d\mathbb{Q}(\mathbf{X}_1, \dots, \mathbf{X}_N)$$

³⁵Fengxiang He, Bohan Wang, and Dacheng Tao. "Tighter Generalization Bounds for Iterative Differentially Private Learning Algorithms". In: 2021.

46 Remarks, Open Questions and Future work on Density Estimation

Optimality : What is the true optimal rate of estimation under ϵ -DP ?

$$\max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\epsilon)^{-\frac{2\beta}{\beta+1}} \right\} \quad \text{VS} \quad \max \left\{ n^{-\frac{2\beta}{2\beta+1}}, (n\epsilon)^{-\frac{2\beta}{\beta+3/2}} \right\}$$

Higher dimensions : What happens when the dimensionality increases ?






47 Remarks, Open Questions and Future work on Quantiles Estimation

Lower bounds : Can we obtain meaningful lower-bounds for the problem ?







"Best of both worlds" mechanism : Can we find a mechanism that behaves like RecExp for a moderate amount of quantiles, but then behaves like Histograms when m is large ?

Functional estimation with regularity assumptions : Can we go from a point estimation problem to a functional estimation one ?

From quantiles to density estimation : We used density estimation to perform quantile function estimation, can we do the converse ?

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